

**Ph.D. Qualifying Exam A**

**Probability Distributions and Regression Analysis**

**January 10, 2023**

1. Let  $Y_1, \dots, Y_5$  be a random sample of size 5 from a standard normal distribution and  $\bar{Y} = \sum_{i=1}^5 Y_i/5$ . Let  $Y_6$  be another observation independent of  $(Y_1, \dots, Y_5)$  also from the standard normal distribution. Prove that  $a\bar{Y} + bY_6$  is independent of the sample variance,  $S^2$  based on  $Y_1, \dots, Y_5$ , where  $a$  and  $b$  are any two fixed real numbers.
2. Let  $X_n$  and  $Y_n$  be random variables such that  $X_n$  and  $Y_n$  are independent for each  $n \geq 1$  and their moment generating functions (m.g.f.'s) exist on the real line. If  $X_n$  converges to  $X$  in distribution and  $Y_n$  converges to  $Y$  in distribution, where  $X$  and  $Y$  are random variables, derive the limiting distribution of  $aX_n + bY_n$ , where  $a$  and  $b$  are any two fixed real numbers.
3. Let  $(X_1, X_2, \dots, X_{k-1})$  have a multinomial distribution with parameters  $(n, p_1, p_2, \dots, p_{k-1})$ , where  $0 < p_i < 1, i = 1, \dots, k, p_1 + p_2 + \dots + p_k = 1$ , and  $X_k = n - (X_1 + X_2 + \dots + X_{k-1})$ .  
(a) Find the m.g.f. of  $(X_2, \dots, X_{k-2})$ . (b) What is the joint probability mass function (p.m.f.) of  $(X_2, X_3, \dots, X_{k-2})$ ? (c) Determine the conditional joint p.m.f. of  $(X_1, X_{k-1})$  given that  $X_2 = x_2, \dots, X_{k-2} = x_{k-2}$ .
4. Consider the normal error linear regression model with two explanatory variables:

$$Y_i = \beta_0 + \beta_1 X_{i1} + \beta_2 X_{i2} + \epsilon_i$$

for  $i \in \{1, 2, \dots, n\}$  and  $\epsilon_i$  are independently identically distributed as  $N(0, \sigma^2)$  distribution.

- (a) Show that the corresponding mean residual sum of squares (EMS) has the expectation

$$E(EMS) = \sigma^2.$$

(b) Show that the corresponding mean regression sum of squares (RMS) has the expectation:

$$E(RMS) = \sigma^2 + \frac{1}{2} \left[ \beta_1^2 \sum_{i=1}^n (X_{i1} - \bar{X}_1)^2 + \beta_2^2 \sum_{i=1}^n (X_{i2} - \bar{X}_2)^2 + 2\beta_1\beta_2 \sum_{i=1}^n (X_{i1} - \bar{X}_1)(X_{i2} - \bar{X}_2) \right].$$

(c) Prove that  $F = RMS/EMS \sim F_{2,n-3}$  under  $H_0 : \beta_1 = \beta_2 = 0$ .

5. Consider the linear model  $Y = X\beta + \epsilon$ , where  $\epsilon$  follows a multivariate normal distribution with 0 mean and diagonal variance-covariance matrix  $\Sigma = \text{diag}(\sigma_1^2, \sigma_2^2, \dots, \sigma_n^2)$ . The vector of weighted-least-squares (WLS) estimates  $\hat{\beta}_w$  can be written as:

$$\hat{\beta}_w = (X'WX)^{-1}X'WY,$$

where  $Y$  is the vector of observed values and  $X$  is the usual model matrix (including a column of 1's) and  $W$  is a diagonal matrix with diagonal elements  $w_1, w_2, \dots, w_n$ .

(a) Suppose there is only one predictor variable  $X$ , derive the explicit formulas for  $\hat{\beta}_{w0}$  and  $\hat{\beta}_{w1}$ .

(b) If  $w_i = 1/\sigma_i^2$  and  $\sigma_i^2$  are proportional to  $X_i$ . Simplify your expressions for  $\hat{\beta}_{w0}$  and  $\hat{\beta}_{w1}$  as much as you can.

6. Consider the multiple linear regression model

$$Y = X\beta + \epsilon \text{ where } \epsilon \sim N(0, \sigma^2 I)$$

Assume that  $\sigma^2$  is unknown. The unknown vector  $\beta$  has dimension  $p \times 1$  with  $p = k + 1$ , and  $k$  is the number of covariates in the model. Consider following partitionings:

$$X = (X_1, X_2), \beta = (\beta_1', \beta_2')'$$

such that

$$X\beta = X_1\beta_1 + X_2\beta_2.$$

Consider the following models:

$$\text{Model 1 : } Y = X_1\beta_1 + X_2\beta_2 + \epsilon$$

and

$$\text{Model 2 : } Y = X_1\beta_1 + \epsilon.$$

Show that  $R_2^2 \leq R_1^2$ , where  $R^2$  is the coefficient of determination.