

**Ph.D. Prelim: Exam A**  
**Distribution Theory & Regression Analysis**

**Jan 2017**

1. Given the cumulative distribution function (c.d.f)

$$F_X(x) = \begin{cases} 0, & x < 0, \\ x^2 + 0.2, & 0 \leq x < 0.5, \\ x, & 0.5 \leq x < 1, \\ 1, & 1 \leq x. \end{cases}$$

- (a) Is  $F_X(x)$  a discrete or continuous distribution?
- (b) Write  $F_X(x)$  in the form of  $cF_1(x) + dF_2(x)$ , where  $F_1$  and  $F_2$  are c.d.f's and  $c$  and  $d$  are known constants. Find  $c$ ,  $d$ ,  $F_1$  and  $F_2$ .
2. (a) Let  $X \geq 0$  be a random variable with finite  $E(X^2)$ . Prove that

$$E(X^2) = \int_0^\infty 2xP(X > x)dx$$

- (b) Let  $X$  and  $Y$  be independent random variables having cumulative distribution functions  $F(t)$  and  $G(t)$  respectively. Suppose that  $(1 - G(t)) = (1 - F(t))^\alpha$ , for all  $t > 0$  and where  $\alpha > 0$ . Prove that  $Z = \min(X, Y)$  and  $\delta = I(X \leq Y)$  are independent random variables.
3. Let  $f$  be a density with support on  $(0, \infty)$ . Let  $g(x, y) = f(x + y)/(x + y)$  for  $x > 0, y > 0$  and 0 otherwise. Prove that  $g$  is a probability density function on  $R^2$  and find its covariance matrix.

4. (a) Prove the following Theorem: If  $y \sim N(0, \sigma^2 I)$  and  $M$  is a symmetric idempotent matrix of rank  $m$ , then

$$\frac{y' M y}{\sigma^2} \sim \chi^2(\text{tr}(M)).$$

- (b) Use above Theorem to show that in a normal error simple linear regression  $Y_i = \beta_0 + X_i \beta_1 + \epsilon_i$  for  $i \in \{1, 2, \dots, n\}$  with  $\epsilon = (\epsilon_1, \epsilon_2, \dots, \epsilon_n)' \sim N(0, \sigma^2 I)$ , the residual sum of squares

$$RSS/\sigma^2 = \sum_{i=1}^n (y_i - \hat{y}_i)^2 / \sigma^2 \sim \chi_{n-2}^2.$$

5. Consider the general linear model  $Y = X\beta + \epsilon$  with  $E(\epsilon) = 0$  and  $\text{cov}(\epsilon) = \sigma^2 I$ , where  $X$  is  $n \times p$  of rank  $p$ . Let  $r_i = Y_i - \hat{Y}_i = Y_i - x_i' \hat{\beta}$  be the  $i$ th residual, where  $\hat{\beta}$  is the least squares estimator of  $\beta$ .

- (a) Show that  $\sum_{i=1}^n \hat{Y}_i (Y_i - \hat{Y}_i) = 0$ .

- (b) Now consider the linear model with an intercept term included explicitly. That is, the first column of  $X$  is all 1's so  $Y_i = \beta_0 + \sum_{j=1}^{p-1} x_{ij}^* \beta_j + \epsilon_i$ ,  $i = 1, 2, \dots, n$  with the same assumptions on the  $\epsilon$ 's as above. Show that in this case  $\sum_{i=1}^n r_i = 0$  and

$$\sum_{i=1}^n (Y_i - \bar{Y})^2 = \sum_{i=1}^n (\hat{Y}_i - \bar{Y})^2 + \sum_{i=1}^n (Y_i - \hat{Y}_i)^2.$$

6. Consider the general linear model  $Y = X\beta + \epsilon$  with  $\epsilon \sim N(0, \sigma^2 I)$ , where  $X$  is  $n \times p$  of rank  $p$ . Let  $\theta = H\beta$  where  $H$  is  $q \times p$  of rank  $q$  and

$$Q = (H\hat{\beta} - \theta)' (H(X'X)^{-1}H')^{-1} (H\hat{\beta} - \theta) / (q \times \hat{\sigma}^2)$$

where  $\hat{\beta}$  is the least squares estimator of  $\beta$ . What is the distribution of  $Q$ ? State what general result you are applying and how it applies here. (You can use the fact that  $(n-p)\hat{\sigma}^2/\sigma^2$  is distributed chi-square with  $n-p$  degrees of freedom and  $\hat{\sigma}^2$  is independent of  $\hat{\beta}$ ).